

**Jakusik Ewa**

*Institute of Meteorology and Water Management - NRI, Warsaw, Poland*

**Kołowrocki Krzysztof**

**Kuligowska Ewa**

**Soszyńska-Budny Joanna**

**Torbicki Mateusz**

*Gdynia Maritime University, Gdynia, Poland*

## **Identification methods and procedures of climate-weather change process including extreme weather hazards of port oil piping transportation system operating at land Baltic seaside area**

### **Keywords**

climate-weather change process, semi-Markov model, modelling, identification, transportation system

### **Abstract**

There are presented the methods of identification of the climate-weather change process. These are the methods and procedures for estimating the unknown basic parameters of the climate-weather change process semi-Markov model and identifying the distributions of the climate-weather change process conditional sojourn times at the climate-weather states. There are given the formulae estimating the probabilities of the climate-weather change process staying at the particular climate-weather states at the initial moment, the probabilities of the climate-weather change transitions between the climate-weather states and the parameters of the distributions suitable and typical for the description of the climate-weather change process conditional sojourn times at the particular climate-weather states. The proposed statistical methods applications for the unknown parameters identification of the climate-weather change process model determining the climate-weather change process parameters for the port oil piping transportation system operating at land Baltic seaside area are presented.

### **1. Introduction**

The general model of the climate-weather change processes is proposed in [3] and [13]. The safety models of various multistate complex technical systems are considered in [5]. Consequently, the general joint models linking these system safety models with the model of their climate-weather processes, allowing us for the safety analysis of the complex technical systems at the variable climate-weather conditions, are constructed in [6]. To be able to apply these general models practically in the evaluation and prediction of the reliability and safety of real complex technical systems it is necessary to have the statistical methods concerned with determining the unknown parameters of the proposed

models [1]-[2], [7]-[8], [10]-[11], [22]. Particularly, concerning the climate-weather process, the probabilities of the climate-weather change process staying at the particular climate-weather states at the initial moment, the probabilities of the climate-weather process transitions between the climate-weather states and the distributions of the conditional sojourn times of the climate-weather process at the particular climate-weather states should be identified [9], [15]-[16]. It is also necessary to use the methods of testing the hypotheses concerned with the climate-weather process conditional sojourn times at the climate-weather states [15].

## 2. Identification and modeling of climate-weather change process

We consider the climate-weather change process  $C(t)$ ,  $t \in \langle 0, +\infty \rangle$ , for the critical infrastructure operating area with  $w$ ,  $w \in N$ , different climate-weather change states  $c_1, c_2, \dots, c_w$  from the set  $\{c_1, c_2, \dots, c_w\}$ . We assume a semi-Markov model [14]-[20], of the climate-weather change process  $C(t)$  and we mark by  $C_{bl}$  its random conditional sojourn times at the climate-weather states  $c_b$ , when its next climate-weather state is  $c_l$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ .

Let  $\Theta$  be the duration time of the experiment. Furthermore, we denote by  $n(0)$  the realisation of the total number of the climate-weather change process stay at the particular climate-weather states at the initial moment  $t = 0$  and by  $[n_b(0)]_{1 \times w}$ ,  $b = 1, 2, \dots, w$ , the vector of realisations of the numbers of staying of the climate-weather change process respectively at the climate-weather states  $c_1, c_2, \dots, c_w$ , at the initial moments  $t = 0$  of all  $n(0)$  observed realizations of the climate-weather change process. Moreover, we denote by  $[n_{bl}]_{w \times w}$  the matrix of the realizations of the numbers  $n_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , of the transitions of the climate-weather change process from the climate-weather state  $c_b$  into the climate-weather state  $c_l$  at all observed realizations of the climate-weather change process. We also denote by  $[n_b]_{1 \times w}$ , the vector of the realizations of the numbers  $n_b$ ,  $b = 1, 2, \dots, w$ , of departures of the climate-weather change process from the climate-weather states  $c_b$ .

Under these assumptions, the climate-weather change process may be described by the vector  $[q_b(0)]_{1 \times w}$  of probabilities of the climate-weather change process staying at the particular climate-weather states at the initial moment  $t = 0$ , the matrix  $[q_{bl}(t)]_{w \times w}$  of the probabilities of the climate-weather change process transitions between the climate-weather states and the matrix  $[C_{bl}(t)]_{w \times w}$  of the distribution functions of the conditional sojourn times  $C_{bl}$  of the climate-weather change process at the climate-weather states or equivalently by the matrix  $[c_{bl}(t)]_{w \times w}$  of the density functions of the conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , of the climate-weather change process at the climate-weather states. These all parameters of the climate-weather change process are unknown and before their use to the prognosis of this process

characteristics have to be estimated on the basis of statistical data coming from practice.

## 3. Statistical identification of climate-weather change process for port oil piping transportation system

### 3.1. Defining parameters and data collection of climate-weather change process for port oil piping transportation system

The unknown parameters of the climate-weather change process semi-Markov model are:

- the initial probabilities  $q_b(0)$ ,  $b = 1, 2, \dots, 36$ , of the climate-weather change process staying at the particular states  $c_b$  at the moment  $t = 0$ ,
- the probabilities  $q_{bl}$ ,  $b, l = 1, 2, \dots, 36$ ,  $b \neq l$ , of the climate-weather change process transitions from the climate-weather state  $c_b$  into the climate-weather state  $c_l$ ,
- the distributions of the climate-weather change conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, 36$ ,  $b \neq l$ , at the particular climate-weather states and their mean values  $M_{bl} = E[C_{bl}]$ ,  $b, l = 1, 2, \dots, 36$ ,  $b \neq l$ .

To identify all these parameters of the climate-weather change process the statistical data about this process is needed.

The collected by the system operators statistical data necessary to evaluating the initial transient probabilities of the climate-weather change process at the particular states are:

- the climate-weather change process observation / experiment time  $\Theta = 8$  years (2007-2015),
- the number of the climate-weather change process realizations  $n(0) = 226$ ,
- the vector of realizations of the numbers of the climate-weather change process staying at the particular climate-weather state  $c_b$  at the initial moment  $t = 0$

$$[n_b(0)] = [9, 1, 0, 0, 0, 0, 0, 89, 13, 0, 0, 0, 0, 0, 144, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0]$$

The collected statistical data necessary to evaluating the probabilities of transitions of the climate-weather state change process  $C(t)$  between the climate-weather states are:

- the matrix of realizations of the numbers of climate-weather change process transitions from the state  $c_b$  into the state  $c_l$  during the experiment time





### 3.4. Identification of distribution functions of climate-weather change process for port oil piping transportation system

Using the procedure given in [4], [16] and the statistical data from Section 3.1 and the results from Section 3.3, we may verify the hypotheses on the distributions of the climate-weather change process conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, 36, b \neq l$ , at the particular states. To do this, we need a sufficient number of realizations of these variables [2], [8], [21], [23]-[24], namely, the sets of their realizations should contain at least 30 realizations coming from the experiment. This condition is not satisfied for the statistical data we have in disposal and that are presented in Section 3.1. However, to make the procedure familiar to the reader, we perform it for the conditional sojourn time  $C_{98}$  having sufficiently numerous set of realizations and preliminarily analyzed in Section 3.3.

The realization  $\bar{c}_{98}(t)$  of the histogram of the climate-weather change process conditional sojourn time  $C_{98}$ , is presented in Table 1 and illustrated in Figure 1

Table 1. The realization of the histogram of the climate-weather change process conditional sojourn time  $C_{98}$

Histogram of the conditional sojourn time $C_{98}$							
$I_j = (a_{98}^j, b_{98}^j)$	0.09 – 1.92	1.92 – 3.75	3.75 – 5.58	5.58 – 7.41	7.41 – 9.24	9.24 – 11.07	11.07 – 12.90
$n_{98}^j$	17	22	4	0	2	0	1
$\bar{h}_{98}(t) = n_{98}^j / n_{98}$	17/46	22/46	4/46	0/46	2/46	0/46	1/46

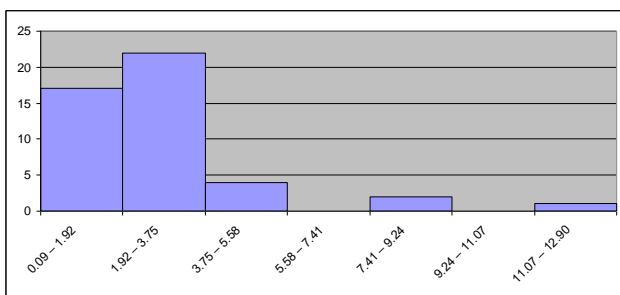


Figure 1. The graph of the histogram of the climate-weather change process conditional sojourn time  $C_{98}$

After analyzing and comparing the realization  $\bar{c}_{98}(t)$  of the histogram with the graphs of the density functions  $c_{bl}(t)$  of the previously distinguished in [3] distributions, we formulate the null hypothesis  $H_0$  in the following form:

$H_0$ : The climate-weather change process conditional sojourn time  $C_{98}$  at the climate-weather state  $c_9$  when the next transition is to the climate-weather state  $c_8$ ,

has the chimney distribution with the density function defined by (4.7) in [3] of the form

$$c_{98}(t) = \begin{cases} 0, & t < x_{98} \\ \frac{A_{98}}{z_{98}^1 - x_{98}}, & x_{98} \leq t \leq z_{98}^1 \\ \frac{K_{98}}{z_{98}^2 - z_{98}^1}, & z_{98}^1 \leq t \leq z_{98}^2 \\ \frac{D_{98}}{y_{98} - z_{98}^2}, & z_{98}^2 \leq t \leq y_{98} \\ 0, & t > y_{98}. \end{cases} \quad (8)$$

We join the intervals defined in the realization of the histogram  $\bar{h}_{98}(t)$  that have the numbers  $n_{98}^j$ , of realizations less than 4 into new intervals and we perform the following steps:

- we fix the new number of intervals  $\bar{r}_{98} = 3$ ,
- we determine the new intervals and we fix the numbers of realizations in the new intervals

Table 2. The numbers of the conditional sojourn time  $C_{98}$  realizations in the intervals  $\bar{I}_j$

$\bar{I}_j = (a_{98}^{-j}, b_{98}^{-j})$	0.09 – 1.92	1.92 – 3.75	3.75 – 12.90
$n_{98}^{-j}$	17	22	7

We estimate the unknown parameters of the density function of the hypothetical chimney distribution using formulae (2.15) - (2.32) in [4] we obtain the following results

$$x_{98} = a_{98}^{-1} = 0.09, \quad y_{98} = x_{98} + \bar{r}_{98} \cdot d_{98} = 12.90,$$

$i = 2$  is the number of the interval including the largest number of realizations i.e. such as that

$$n_{98}^{-2} = \max_{1 \leq j \leq \bar{r}_{98}} \{ n_{98}^{-j} \} = 22,$$

$$n_{98}^{-1} = 17 \neq 0, \quad \frac{n_{98}^{-2}}{n_{98}^{-1}} = \frac{22}{17} < 3 \quad \text{and} \quad \frac{n_{98}^{-2}}{n_{98}^{-3}} = \frac{22}{7} \geq 3,$$

so

$$z_{98}^1 = x_{98} + (i - 2)d_{98} = 0.09,$$

$$z_{98}^2 = x_{98} + id_{98} = 3.75, \quad A_{98} = \frac{0}{n_{98}} = 0,$$

$$K_{89} = \frac{n_{98}^{-i-1} + n_{98}^{-i}}{n_{98}} = 0.85, \quad D_{98} = \frac{n_{98}^{-i+1}}{n_{98}} = 0.15.$$

- we calculate the hypothetical probabilities that the variable  $C_{98}$  takes values from the new intervals

$$\begin{aligned} p_1 &= P(C_{98} \in \bar{I}_1) = P(0.09 \leq C_{98} < 1.92) \\ &= C_{98}(1.92) - C_{98}(0.09) = 0.43 - 0 = 0.43, \\ p_2 &= P(C_{98} \in \bar{I}_2) = P(1.92 \leq C_{98} < 3.75) \\ &= C_{98}(3.75) - C_{98}(1.92) = 0.86 - 0.43 = 0.43, \\ p_3 &= P(C_{98} \in \bar{I}_3) = P(3.75 \leq C_{98} < 12.90) \\ &= C_{98}(12.90) - C_{98}(3.75) = 1 - 0.86 = 0.14, \end{aligned}$$

- we calculate the realization of the  $\chi^2$  (chi-square)-Pearson's statistics

$$\begin{aligned} u_{98} &= \sum_{j=1}^3 \frac{(\bar{n}_{98}^j - n_{98} \cdot p_j)^2}{n_{98} \cdot p_j} = \frac{(17 - 46 \cdot 0.43)^2}{46 \cdot 0.43} \\ &\quad + \frac{(22 - 46 \cdot 0.43)^2}{46 \cdot 0.43} + \frac{(7 - 46 \cdot 0.14)^2}{46 \cdot 0.14} \\ &\cong 0.3907 + 0.2492 + 0.0487 = 0.6886 \cong 0.69, \end{aligned}$$

- we assume the significance level  $\alpha = 0.05$ ,  
 - we fix the number of degrees of freedom

$$\bar{r}_{98} - l - 1 = 3 - 0 - 1 = 2,$$

- we read from the tables of the  $\chi^2$  - Pearson's distribution the value  $u_\alpha$  for the fixed values of the significance level  $\alpha = 0.05$  and the number of degrees of freedom  $\bar{r}_{98} - l - 1 = 2$ , such that, according to (2.48), the following equality holds

$$P(U_{98} > u_\alpha) = \alpha = 0.05$$

that amounts  $u_\alpha = 5.99$  and we determine the critical domain in the form of the interval  $(5.99, +\infty)$  and the acceptance domain in the form of the interval  $< 0, 5.99 >$ .

- we compare the obtained value  $u_{98} = 0.69$  of the realization of the statistics  $U_{98}$  with the read from the tables critical value  $u_\alpha = 5.99$  of the chi-square random variable and since the value  $u_{98} = 0.69$  does not belong to the critical domain, i.e.

$$u_{98} = 0.69 \leq u_\alpha = 5.99,$$

then we do not reject the hypothesis  $H_0$ , that the sojourn time  $C_{98}$  has the chimney distribution with the density function given by (8).

For the remaining cases, when the realizations of conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, 36$ ,  $b \neq l$ , at the particular climate-weather states are more than

30, proceeding afterwards in an analogous way as in the case of the conditional sojourn time  $C_{98}$ , we can get the following results:

- the climate-weather change process conditional sojourn time  $C_{89}$  has the exponential distribution with the density function defined by (4.5) in [3] of the form

$$c_{89}(t) = \begin{cases} 0, & t < x_{89} \\ \alpha_{89} \exp[-\alpha_{89}(t - x_{89})], & t \geq x_{89}, \end{cases}$$

with the parameters

$$\begin{aligned} x_{89} &= a_{89}^1 = 0, \\ \alpha_{89} &= \frac{1}{\bar{C}_{89} - x_{89}} = \frac{1}{14.40 - 0} \cong 0.0694. \end{aligned}$$

- the climate-weather change process conditional sojourn time  $C_{814}$  has the exponential distribution with the density function defined by (4.5) in [3] of the form

$$c_{814}(t) = \begin{cases} 0, & t < x_{814} \\ \alpha_{814} \exp[-\alpha_{814}(t - x_{814})], & t \geq x_{814}, \end{cases}$$

with the parameters

$$\begin{aligned} x_{814} &= a_{814}^1 = 0, \\ \alpha_{814} &= \frac{1}{\bar{C}_{814} - x_{814}} = \frac{1}{31.65 - 0} \cong 0.0316. \end{aligned}$$

- the climate-weather change process conditional sojourn time  $C_{915}$  has the exponential distribution with the density function defined by (4.5) in [3] of the form

$$c_{915}(t) = \begin{cases} 0, & t < x_{915} \\ \alpha_{915} \exp[-\alpha_{915}(t - x_{915})], & t \geq x_{915}, \end{cases}$$

with the parameters

$$\begin{aligned} x_{915} &= a_{915}^1 = 0, \\ \alpha_{915} &= \frac{1}{\bar{C}_{915} - x_{915}} = \frac{1}{2.95 - 0} \cong 0.3390. \end{aligned}$$

- the climate-weather change process conditional sojourn time  $C_{159}$  has the exponential distribution with the density function defined by (4.5) in [3] of the form

$$c_{159}(t) = \begin{cases} 0, & t < x_{159} \\ \alpha_{159} \exp[-\alpha_{159}(t - x_{159})], & t \geq x_{159}, \end{cases}$$

with the parameters

$$x_{159} = a_{159}^1 = 0, \\ \alpha_{159} = \frac{1}{\bar{C}_{159} - x_{159}} = \frac{1}{46.29 - 0} \cong 0.0216.$$

- the climate-weather change process conditional sojourn time  $C_{15\ 21}$  has the exponential distribution with the density function defined by (4.5) in [3] of the form

$$c_{15\ 21}(t) = \begin{cases} 0, & t < x_{15\ 21} \\ \alpha_{15\ 21} \exp[-\alpha_{15\ 21}(t - x_{15\ 21})], & t \geq x_{15\ 21}, \end{cases}$$

with the parameters

$$x_{15\ 21} = a_{15\ 21}^1 = 0, \\ \alpha_{15\ 21} = \frac{1}{\bar{C}_{15\ 21} - x_{15\ 21}} = \frac{1}{60.25 - 0} \cong 0.0166.$$

For the distributions identified in this section, by application either the general formulae for the mean value given by (2.12) or the particular formulae (2.13)-(2.19) in [16], the mean values  $M_{bl} = E[\theta_{bl}]$ ,  $b, l = 1, 2, \dots, 36$ ,  $b \neq l$ , of the port oil piping transportation climate-weather change process conditional sojourn times at the particular climate-weather states can be determined and they amount:

$$M_{89} \cong 14.41, M_{8\ 14} \cong 31.65, M_{98} \cong 2.88. \\ M_{9\ 15} \cong 2.95, M_{159} \cong 46.30, M_{15\ 21} \cong 60.24.$$

Because of the lack of sufficient numbers of realizations of the climate-weather change process conditional sojourn times at the climate-weather states, it is not possible to identify statistically their distributions. In those cases of not identified distributions it is possible to find the approximate empirical values of the mean values  $M_{bl} = E[C_{bl}]$  of the conditional sojourn times at the particular climate-weather states that are as follows:

$$M_{12} = 9.50, M_{17} = 18.25, M_{18} = 4.40, M_{21} = 1.80, \\ M_{28} = 2.00, M_{71} = 1.00, M_{78} = 1.50, M_{81} \cong 69.67, \\ M_{82} \cong 19.67, M_{8\ 15} \cong 26.29, M_{14\ 8} \cong 2.19, \\ M_{14\ 9} = 2.00, M_{14\ 15} = 1.75, M_{15\ 8} = 40.05, \\ M_{15\ 14} = 1.00, M_{15\ 16} = 35.25, M_{16\ 15} = 2.00,$$

$$M_{16\ 22} = 1.50, M_{21\ 15} = 1.82, M_{21\ 22} = 2.00, \\ M_{22\ 15} = 1.00, M_{22\ 16} \cong 1.67, M_{22\ 21} = 1.00.$$

As there are no realizations for the rest conditional sojourn times at the climate-weather states of the port oil piping transportation climate-weather change process, it is impossible to estimate their empirical conditional mean values.

#### 4. Conclusions

The proposed statistical methods of identification of the unknown parameters of the climate-weather change processes allow us for the identification of the models discussed in [6] and next their practical applications in evaluation, prediction and optimization of reliability, availability and safety of real complex critical infrastructures.

#### Acknowledgments



The paper presents the results developed in the scope of the EU-CIRCLE project titled "A pan – European framework for strengthening Critical Infrastructure resilience to climate change" that has received funding from the European Union's Horizon 2020 research and innovation programme under grant agreement No 653824. <http://www.eu-circle.eu/>.

#### References

- [1] Barbu, V., Limnios, N. (2006). Empirical estimation for discrete-time semi-Markov processes with applications in reliability. *Journal of Nonparametric Statistics* 18, 7-8, 483-498.
- [2] Collet, J. (1996). Some remarks on rare-event approximation. *IEEE Transactions on Reliability* 45, 106-108.
- [3] EU-CIRCLE Report D2.1-GMU3. (2016). *Modelling Climate-Weather Change Process Including Extreme Weather Hazards*.
- [4] EU-CIRCLE Report D2.3-GMU2. (2016). *Identification Methods and Procedures of Climate-Weather Change Process Including Extreme Weather Hazards*.
- [5] EU-CIRCLE Report D3.3-GMU1. (2016). *Modelling inside dependences influence on safety of multistate ageing systems – Modelling safety of multistate ageing systems*.
- [6] EU-CIRCLE Report D3.3-GMU12. (2017). *Integration of the Integrated Model of Critical Infrastructure Safety (IMCIS) and the Critical Infrastructure Operation Process General Model (CIOPGM) into the General Integrated Model of Critical Infrastructure Safety (GIMCIS) related to*

- operating environment threads (OET) and climate-weather extreme hazards (EWH).
- [7] Gamiz, M. L. & Roman, Y. (2008). Non-parametric estimation of the availability in a general repairable. *Reliability Engineering & System Safety* 93, 8, 1188-1196.
- [8] Giudici, P. & Figini, S. (2009). Applied data mining for business and industry. John Wiley & Sons Ltd.
- [9] Habibullah, M. S., Lumanpauw, E., Kołowrocki, K. et al. (2009). A computational tool for general model of operation processes in industrial systems operation processes. *Electronic Journal Reliability & Risk Analysis: Theory & Applications* 2, 4, 181-191.
- [10] Helvacioğlu, S. & Insel, M. (2008). Expert system applications in marine technologies. *Ocean Engineering* 35, 11-12, 1067-1074.
- [11] Hryniewicz, O. (1995). Lifetime tests for imprecise data and fuzzy reliability requirements. *Reliability and Safety Analyses under Fuzziness*. Onisawa, T. & Kacprzyk, J. (Ed.). Physica Verlag, Heidelberg, 169-182.
- [12] Jakusik, E., Kołowrocki, K., Kuligowska, E. et al. (2016). Identification methods and procedures of climate-weather change process including extreme weather hazards for oil piping transportation system operating at under water Baltic Sea area, *Summer Safety & Reliability Seminars. Journal of Polish Safety and Reliability Association* 7, 3, 47-56.
- [13] Jakusik, E., Kołowrocki, K., Kuligowska, E. et al. (2016). Modelling climate-weather change process including extreme weather hazards for oil piping transportation system, *Summer Safety & Reliability Seminars. Journal of Polish Safety and Reliability Association* 7, 3, 31-40.
- [14] Kołowrocki, K. (2004). Reliability of Large Systems, Elsevier, ISBN: 0080444296.
- [15] Kołowrocki, K. (2014). Reliability of large and complex systems, Elsevier, ISBN: 978080999494.
- [16] Kołowrocki, K. & Soszyńska-Budny, J. (2011). Reliability and Safety of Complex Technical Systems and Processes: Modeling-Identification-Prediction-Optimization. Springer, ISBN: 9780857296931.
- [17] Limnios, N. & Oprisan, G. (2005). Semi-Markov Processes and Reliability. Birkhauser, Boston.
- [18] Limnios, N., Ouhbi, B. & Sadek, A. (2005). Empirical estimator of stationary distribution for semi-Markov processes. *Communications in Statistics-Theory and Methods* 34, 4, 987-995.
- [19] Macci, C. (2008). Large deviations for empirical estimators of the stationary distribution of a semi-Markov process with finite state space. *Communications in Statistics-Theory and Methods* 37, 19, 3077-3089.
- [20] Mercier, S. (2008). Numerical bounds for semi-Markovian quantities and application to reliability. *Methodology and Computing in Applied Probability* 10, 2, 179-198.
- [21] Rice, J. A. (2007). Mathematical statistics and data analysis. Duxbury. Thomson Brooks/Cole. University of California. Berkeley.
- [22] Soszyńska, J., Kołowrocki, K., Blokus-Roszkowska, A. et al. (2010). Identification of complex technical system components safety models. *Summer Safety & Reliability Seminars. Journal of Polish Safety and Reliability Association* 42, 399-496.
- [23] Vercellis, S. (2009). Data mining and optimization for decision making. John Wiley & Sons Ltd.
- [24] Wilson, A. G., Graves, T. L., Hamada, M. S. et al (2006). Advances in data combination, analysis and collection for system reliability assessment. *Statistical Science* 21, 4, 514-531.